

Final Exam in System Identification for F and STS

Answers and Brief Solutions

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1. a)

$$\lim_{N \rightarrow \infty} \hat{\theta}_N - \theta_0 = \begin{bmatrix} \text{E}\{y^2(t)\} & -\text{E}\{y(t)u(t)\} \\ -\text{E}\{u(t)y(t)\} & \text{E}\{u^2(t)\} \end{bmatrix}^{-1} \begin{bmatrix} -\text{E}\{y(t-1)w(t)\} \\ \text{E}\{u(t-1)w(t)\} \end{bmatrix} \neq 0$$

b) For example,

$$\check{\theta}_N = \left(\frac{1}{N} \sum_{t=1}^N z(t)\varphi^T(t) \right)^{-1} \left(\frac{1}{N} \sum_{t=1}^N z(t)y(t) \right)$$

where

$$z(t) = [-y(t-2) \quad u(t-1)]^T$$

2. a)

$$\text{mse}(\check{d}) = \text{var}(\check{d}) + \underbrace{(\text{bias}(\check{d}))^2}_{=0} = \text{var}(\check{d}) = \frac{\lambda^2}{N}$$

b)

$$\text{mse}(\bar{d}) = \text{var}(\bar{d}) + (\text{bias}(\bar{d}))^2 = \frac{a^2\lambda^2}{N} + d^2(a-1)^2$$

where

$$a_{\text{opt}} = \frac{d^2}{d^2 + \lambda^2/N}$$

minimizes $\text{mse}(\bar{d})$. The estimator \bar{d} is not realizable with this choice of a , since a_{opt} depends on the unknown parameter d .

c)

$$\hat{d}(t) = \left(\sum_{k=1}^t \lambda^{t-k} \right)^{-1} \sum_{k=1}^t \lambda^{t-k} x(k)$$

$$\hat{d}(t) = \hat{d}(t-1) + \frac{1-\lambda}{1-\lambda^t} (x(t) - \hat{d}(t-1))$$

3. a) Not identifiable; predictor $\hat{y}(t|b_1, b_2) = (b_1 + b_2)u_0$.
- b) Not identifiable; predictor $\hat{y}(t|a, b) = (-bc - a)y(t - 1)$.
Use, for example, $u(t) = -cy(t - 1)$ or use two different values of c .
4. a) One resonance peak is given by a complex conjugated pair of poles. Start with model order four.
- b) ARX; linear regression.
- c) Overfit.
- d) The result is often given as a table or as a curve.
- e) The A -polynomial of the ARX-model must describe the disturbance dynamic through $1/A$. This can result in a slightly erroneous description of the system dynamic. It is easier for the ARMAX-model to describe both the system- and the disturbance dynamic due to the C -polynomial.
- f) A useful and good model can describe new data with high enough accuracy. Such a test can also reduce the risk for overfit.

5. Predictor

$$\hat{y}(t|\theta) = [-y(t-1) \quad -y(t-2)] \begin{bmatrix} a_1 \\ a_2 \end{bmatrix} = \varphi^T(t)\theta$$

a)

$$\begin{aligned} \mathbf{E}\{(\hat{\theta}_N - \theta_0)(\hat{\theta}_N - \theta_0)^T\} &\approx \frac{\lambda^2}{N} (\mathbf{E}\{\varphi(t)\varphi^T(t)\})^{-1} \\ &= \frac{\lambda^2}{N} \begin{bmatrix} r_y(0) & r_y(1) \\ r_y(1) & r_y(0) \end{bmatrix}^{-1} = \frac{1}{N} \begin{bmatrix} 1 & a_0 \\ a_0 & 1 \end{bmatrix} \end{aligned}$$

so $\text{var}(\hat{a}_1) = \text{var}(\hat{a}_2) = 1/N$.

b)

$$\begin{aligned} \lim_{N \rightarrow \infty} \hat{\theta}_N &= (\mathbf{E}\{\varphi(t)\varphi^T(t)\})^{-1} \mathbf{E}\{\varphi(t)y(t)\} \\ &= \begin{bmatrix} r_y(0) & r_y(1) \\ r_y(1) & r_y(0) \end{bmatrix}^{-1} \begin{bmatrix} -r_y(1) \\ -r_y(2) \end{bmatrix} = \begin{bmatrix} a_0 \\ 0 \end{bmatrix} \end{aligned}$$

$\lim_{N \rightarrow \infty} \hat{a}_1 = a_0$. The estimate is correct.

$\lim_{N \rightarrow \infty} \hat{a}_2 = 0$. This makes sense when comparing the model with the system.

6. a)

$$\hat{G}(e^{i\omega}) = \sum_{k=0}^{\infty} \hat{g}(k)e^{-i\omega k} = \frac{1}{\alpha} \sum_{k=0}^{\infty} y(k)e^{-i\omega k} = \frac{1}{\alpha} \sqrt{N} Y_N(\omega) = \frac{Y_N(\omega)}{U_N(\omega)} = \hat{G}_N(e^{i\omega})$$

where it is used that

$$U_N(\omega) = \frac{1}{\sqrt{N}}\alpha$$

in the second last equality.

b)

$$\begin{aligned}\tilde{g}(t) &= \hat{g}(t) - g_0(t) = \frac{v(t) - v(t-1)}{\beta} \\ \text{var}(\tilde{g}(t)) &= \frac{2\lambda^2}{\beta^2}\end{aligned}$$

7. Linear regression

$$-\hat{r}(\tau) = [\hat{r}(\tau-1) \quad \cdots \quad \hat{r}(\tau-na)] \begin{bmatrix} a_1 \\ \vdots \\ a_{na} \end{bmatrix}$$

where

$$\hat{r}(\tau) = \frac{1}{N} \sum_{t=\tau}^N y(t)y(t-\tau)$$

$\tau = nc+1, nc+2, \dots, nc+na$ give

$$\begin{bmatrix} \hat{r}(nc) & \cdots & \hat{r}(nc+1-na) \\ \vdots & & \vdots \\ \hat{r}(nc+na-1) & \cdots & \hat{r}(nc) \end{bmatrix} \begin{bmatrix} \hat{a}_1 \\ \vdots \\ \hat{a}_{na} \end{bmatrix} = \begin{bmatrix} -\hat{r}(nc+1) \\ \vdots \\ -\hat{r}(nc+na) \end{bmatrix}$$

This can (essentially, apart from different start indexes in sums) be written as

$$\left(\frac{1}{N} \sum_{t=1}^N z(t)\varphi^T(t) \right) \hat{\theta}_N^{\text{IV}} = \frac{1}{N} \sum_{t=1}^N z(t)y(t)$$

where

$$\varphi(t) = [-y(t-1) \quad \cdots \quad -y(t-na)]^T$$

(the AR-part is estimated) and

$$z(t) = [-y(t-(nc+1)) \quad \cdots \quad -y(t-(nc+na))]^T$$

8. a) Let

$$\hat{\lambda}_N = \frac{1}{N} \sum_{t=1}^N \varepsilon^2(t, \hat{\theta}_N^{(p)})$$

The C_p -criterion can then be written as

$$(1) \quad C_p = \frac{N\hat{\lambda}_N}{\hat{s}_N^2} - (N - 2p)$$

When selecting p , the quantities N and

$$\hat{s}_N^2 = \frac{1}{N} \sum_{t=1}^N \varepsilon^2(t, \hat{\theta}_N^{(p_{\max})})$$

are seen as constants. Minimizing (1) with respect to p is therefore equivalent to minimizing

$$(2) \quad N\hat{\lambda}_N + 2p\hat{s}_N^2$$

with respect to p . Compare with AIC, which minimizes

$$(3) \quad N\hat{\lambda}_N \left(1 + \frac{2p}{N}\right) = N\hat{\lambda}_N + 2p\hat{\lambda}_N$$

The only difference between (2) and (3) is that $\hat{\lambda}_N$ in (3) is replaced by \hat{s}_N^2 . From an operational point of view, this difference is minor.

b) Assume that the denominator is

$$C(z) = (z - p_1)(z - p_2)$$

so

$$C(e^{i\omega}) = (e^{i\omega} - p_1)(e^{i\omega} - p_2)$$

Resonance peaks at $\omega = \pm 1$ means that $C(e^{i\omega})$ has minimum values for $\omega = \pm 1$:

$$\begin{aligned} e^i - p_1 &= 0 \\ e^{-i} - p_2 &= 0 \end{aligned}$$

so $p_1 = e^i$ and $p_2 = e^{-i}$.

c) Minimum point of $Q(x)$? Solve $Q'(x) = 0$ to get

$$x = x_k - (J''(x_k))^{-1} J'(x_k)$$

Take x_{k+1} as the minimum point:

$$x_{k+1} = x_k - (J''(x_k))^{-1} J'(x_k)$$