System Identification, Exercises

Kristiaan Pelckmans (IT/UU, 2338)

Course code: 1RT875, Report code: 61806, F, FRI Uppsala University, Information Technology

January - March 2010

Requirements Exam

- Exercises alike and working knowledge.
- Required (print assignments, add solution pages, put name/option and hand in at lecture)
 - Answers to computer exercises.
 - Homework assignments.
 - Answers to laboratory session.
- Handed out during respective labs.
- Deadline end lectures (1 march)
- Written exam ± 18 march.
- Ph.D. please contact me = kp@it.uu.se.

Book

- Introduction (Chap.2)
- Nonparametric Methods (Chap.3)
- Linear Regression (Chap.4)
- Input Signals (Chap.5)
- Model Parametrizations (Chap.6)
- Prediction Error Methods (Chap.7)
- Instrumental Variable Methods (Chap.8)
- Recursive Identification (Chap.9)
- Identification of Systems Operating in Closed Loop (Chap.10)
- Model Validation (Chap.11)

Key formulas

- Deterministic vs. stochastic.
- Expectation (for ergodic, stationary timeseries $(y(t))_t$)

$$E[f(y(t))] = \frac{1}{N} \sum_{t=1}^{N} f(y(t))$$

- $(y(t))_t$ and $(e(t))_t$ independent timeseries iff $\forall f, g$ E[f(y(t))g(e(t))] = E[f(y(t))]E[g(e(t))]
- Bias $\theta_0 E[\hat{\theta}_N]$, consistent if $\lim_{N \to \infty} \hat{\theta}_N = \theta_0$
- System, model, parametrisation, estimator.
- Covariance of an estimate

$$\operatorname{cov}(\hat{\theta}_N) = E\left[(\hat{\theta}_N - \theta_0)^T(\hat{\theta}_N - \theta_0)\right]$$

- Least Squares estimator
- PE, SI, PI

Given a system

$$H_1(z) = \frac{b}{z+a}, \quad H_1(z) = \frac{b_0 z + b_1}{z^2 + a_1 z + a_2}$$

1. If this filters white noise, zero mean, unit variance and

$$\phi_y(\omega) = \frac{1}{2\pi} \frac{0.75}{1.25 - \cos(\omega)}.$$

What is the variance of the filtered signal?

- 2. What happens to the output of the second system when you move the poles of $H_2(z)$ towards the unit circle?
- 3. Where to place the poles to get a 'low-pass' filter?
- 4. Where to put the poles in order to have a resonance top at $\omega = 1$?

- 5. How does a resonant system appear on the different plots?
- 6. What happens if $H_2(z)$ got a zero close to the unit circle?

Determine the covariance function for an AR(1) process

$$y(t) + a(y(t-1)) = e(t)$$

where e(t) white, zero mean and unit variance.

Determine the covariance function for a MA(1) process

$$y(t) = e(t) + ce(t-1)$$

where e(t) white, zero mean and unit variance.

Consider a general MA(n). For which values τ is it in general true that $r(\tau)=0?$

Given an input u(t) shaped by an ARMA filter,

$$A(q^{-1})x(t) = C(q^{-1})v(t)$$

where v(t) white, zero mean and variance $\lambda_v^2.\,$ Given noisy observations of this signal, or

$$y(t) = x(t) + e(t)$$

where e(t) white, zero mean and variance λ_e^2 and uncorrelated to v(t). Rewrite this as a ARMA process, what would be the corresponding variance of the 'noise'? How would the spectrum of y(t) look like?

Convergence rates for consistent estimators.

For most consistent estimators of the parameters of stationary processes, the estimation error $\hat{\theta} - \theta_0$ tends to zero as 1/N when $N \to \infty$. For nonstationary processes, faster convergence rates may be expected. To see this, derive the variance of the least squares estimate in the model

$$y(t) = \alpha t + e(t), \ t = 1, \dots, N$$

with e(t) white noise, zero mean and variance λ^2 .

Illustration of unbiasedness and consistency properties. Let $\{x_i\}_i$ be a sequence of i.i.d. Gaussian random variables with mean μ and variance σ . Both are unknown. Consider the following estimate of μ :

$$\hat{\mu} = \frac{1}{N} \sum_{i=1}^{N} x_i$$

and the following two estimates of σ :

$$\begin{cases} \hat{\sigma}_1 = \frac{1}{N} \sum_{i=1}^N (x_i - \hat{\mu})^2 \\ \hat{\sigma}_2 = \frac{1}{N-1} \sum_{i=1}^N (x_i - \hat{\mu})^2 \end{cases}$$

determine the mean and the variance of the estimates $\hat{\mu}, \hat{\sigma}_1$ and $\hat{\sigma}_2$. Discuss their bias and consistency properties. Compare $\hat{\sigma}_1$ and $\hat{\sigma}_2$ in terms of their Mean Square Error (mse).

SI-2010

Least square estimates with white noise as inputs.

Given a system

$$y(t) + a_0 y(t-1) = b_0 u(t-1) + e(t) + c_0 e(t-1)$$

with e(t) white, zero mean and variance λ^2 , and zero mean white noise input $(u(t))_t$ with variance σ^2 , uncorrelated with noise $e(s), s \leq t$, then

$$\begin{cases} E[y(t)u(t)] = 0\\ E[y(t)y(t-1)] = \frac{-a_0b_0^2 + (c_0 - a_0)(1 - a_0c_0)\lambda^2}{(1 - a_0^2)}\\ E[y(t)u(t-1)] = b_0\sigma^2 \end{cases}$$

rewrite as a model LIP:

$$y(t) = (y(t-1), u(t-1))^T(\alpha, \beta)$$

Application of LS yields estimates for $N \to \infty$

$$\begin{cases} \hat{\alpha} = a_0 + \frac{-c_0(1-a_0^2)\lambda^2}{b_0^2 \sigma^2 + (1+c_0^2 - 2a_0 c_0)\lambda^2} \\ \hat{\beta} = b_0 \end{cases}$$

Least square estimates of the same system with a step function $u(t)=\sigma I(t>0)$ as inputs.

Let $S = \frac{b_0}{1+a_0}$. The covariance matrices become

$$\begin{cases} E[y^{2}(t)] = S^{2}\sigma^{2} + \frac{(1+c_{0}^{2}-2a_{0}c_{0})\lambda^{2}}{1-a_{0}^{2}} \\ E[u^{2}(t)] = \sigma^{2} \\ E[y(t)u(t)] = S\sigma^{2} \\ E[y(t)u(t)] = S\sigma^{2} \\ E[y(t)u(t-1)] = S\sigma^{2} \\ E[y(t)y(t-1)] = S^{2}\sigma^{2} + \frac{(c_{0}-a_{0})(1-a_{0}c_{0})\lambda^{2}}{1-a_{0}^{2}} \end{cases}$$

Application of LS yields estimates for $N \to \infty$

$$\begin{cases} \hat{\alpha} = a_0 - \frac{c_0(1-a_0^2)}{(1+c_0^2 - 2a_0c_0)} \\ \hat{\beta} = b_0 - b_0c_0 \left(\frac{1-a_0}{1+c_0^2 - 2a_0c_0}\right) \end{cases}$$

K. Pelckmans

Least square estimates of the same system with a step function $u(t) = \sigma I(t > 0)$ as inputs (Ct'd).

Verify that the gain S is estimated correctly when $N \to \infty$ by

$$\hat{S} = \frac{\hat{\beta}}{1+\hat{\alpha}} = \frac{b_0}{1+a_0}$$

In the noise-free case where $\lambda = 0$, we run into troubles, the covariance matrix is noninvertible:

$$\begin{bmatrix} E[y^2(t)] & E[y(t)u(t)] \\ E[u(t)y(t)] & E[u^2(t)] \end{bmatrix}$$

How are all possible solutions characterized?

Determine the time constant T from a step response. A first order system Y(s) = G(s)U(s) with

$$G(s) = \frac{K}{1+sT}e^{-s\tau}$$

or in time domain as a differential equation

$$T\frac{dy(t)}{dt} + y(t) = Ku(t - \tau)$$

derive a formula of the step response of an input u(t) = I(t > 0).

Correlation analysis with truncated weighting function.

$$\begin{bmatrix} r_{uy}(0) \\ \vdots \\ r_{uy}(M-1) \end{bmatrix} \begin{bmatrix} r_u(0) & r_u(M-1) \\ \vdots \\ r_u(M-1) & r_{uy}(0) \end{bmatrix} \begin{bmatrix} \hat{h}(0) \\ \vdots \\ \hat{h}(M-1) \end{bmatrix}$$

- 1. Let an input u(t) be white noise, note that regardless of M the solution $\hat{h}(k) = h_0(k)$ for $k = 0, \dots, M 1$.
- 2. Consider the input ($|\alpha| < 1$, and v(t) zero mean white noise with variance σ^2).

$$u(t) - \alpha u(t-1) = v(t)$$

and assume a first order system

$$y(t) + ay(t-1) = bu(t-1), |a| < 1$$

Then

$$\begin{cases} h(0) = 0\\ h(k) = b(-a)^{k-1}, \ \forall k \ge 1\\ \hat{h}(k) = h_0(k), \ k = 0, \dots, M - 2\hat{h}(M-1) = \frac{h(M-1)}{1+a\alpha} \end{cases}$$

Hint:

$$\begin{bmatrix} 1 & \alpha & \dots & \alpha^{M-1} \\ \alpha & 1 & & \\ & \ddots & & \\ \alpha^{M-1} & & & 1 \end{bmatrix} = \frac{1}{1-\alpha^2} \begin{bmatrix} 1 & -\alpha & & \\ -\alpha & 1+\alpha^2 & & \\ & & \ddots & \\ & & & \ddots & \\ & & & & 1 \end{bmatrix}$$

Step response as a special case of spectral analysis.

Let $(y(t))_t$ be the step response of an LTI $H(q^{-1})$ to an input $u(t) = aI(t \ge 1)$. Assume y(t) = 0 for t < 1 and $y(t) \approx c$ for t > N. Justify the following rough estimate of H

$$\hat{h}(k) = \frac{y(k) - y(k-1)}{a}, \ \forall k = 0, \dots, N$$

and show that it is approximatively equal to the estimate provided by the spectral analysis.